Generalized options, utilities and information measures

Reading circle on Bayesian theory, October 10th 2006

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Outline

Exercise from last week

Generalized options and utilities

Generalized information measures

Intrinsic Estimation

Exercise from last week

Find the expectation of $f(x) = \frac{1}{1+x^2}$

Answer:

Exercise from last week

Find the expectation of $f(x) = \frac{1}{1+x^2}$

Answer:

•
$$E[x] = \int_{-\infty}^{\infty} \frac{x}{1+x^2} \mathrm{d}x.$$

• Now, let
$$F(x) = \ln(1 + x^2)$$
, in which case $\int_a^b \frac{x}{1+x^2} dx = \frac{F(b)}{2} - \frac{F(b)}{2}$

► Here $a = -\infty$ and $b = \infty$ and thus $E[x] = \frac{F(\infty)}{2} - \frac{F(-\infty)}{2}$, which is not defined

Generalized options and utilities

Motivation and preliminaries

- The chapter 3.2 extended the quantitative coherence theory (chapters 2.1-2.4) to the infinite domain
- The chapters 3.3 and 3.4 generalize the theory of options, utilities and information measure (chapters 2.5-2.7) to the infinite domain

Generalized options and utilities

Motivation and preliminaries

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key features are

- Convergence in expected utility
- Definition of decision
- propositions that ensure the existence of sequencies of options a that converge in expected utility to the decision
 - needed in defining (finding) the decision

Generalized options and utilities

Generalized preferences

- Extension of the preference relation
- Maximization of expected utility for decision
 - also in the case of conditional

The value of information

- Optimal experimental design
- The value of additional information
- Expected value of perfect information
 - Additive decomposition

Generalized information measures

The utility of general probability distribution

- Score function
 - mapping $u \ \mathfrak{Q} \times \Omega \to \mathfrak{R}$
- Prober score function
 - if expected score is maximized when $q_{\omega}(.|D) = p_{\omega}(.|D)$

Generalized information measures

The utility of general probability distribution

- Score function
 - mapping $u \ \mathfrak{Q} \times \Omega \to \mathfrak{R}$
- Prober score function
 - if expected score is maximized when $q_{\omega}(.|D) = p_{\omega}(.|D)$
- quadratic score function
 - Proved to be proper
- Local score function
 - *u* is local if there exist functions *u*_∞ such that *u*(*q*_∞(.|*D*)) = *u*_∞(*q*(∞|*D*))
- ► Proper local score functions are logarithmic score function Alog (q(ω|D) + B(ω))

Generalized information measures

Generalized approximation and discrepancy

- Expected loss in probability reporting
 - Based on the logarithmic score function
- Discrepancy of an approximation
 - the expected loss in approximating p with \hat{p}

Generalized information

- Information from data
 - Is the expected utility of data (given the prior information) in the sense of logarithmic score function
- Expected information from an experiment

- The final result of Bayesian inference is the posterior distribution of the quantity of interest
- However, often it is necessary to be able to give a good point estimate
 - Good estimate should be objective and invariant under one-to-one transformations
 - How to find a good estimate?
- The problem can be formulated as a decision problem
 - let p(x |θ) be a probability model assumed to describe data x,
 - ► now action space is A = {θ^e}, where θ^e is possible point estimate

Loss function

- ► Let $I(\theta^e, \theta^a)$ be a loss function measuring the consecuence of estimating θ^a (the actual true parameter value) with θ^e
- conventional loss functions compare θ^e to θ^a
 - for example squared, zero-one and absolute error loss
 - Problems in invariance under one-to-one transformations and generalization into higher dimensions
- Intrinsic loss function compares $p(\mathbf{x} | \theta^e)$ to $p(\mathbf{x} | \theta^a)$
 - for example, the one based on logarithmic score function mentioned earlier (Kullback-Leibler divergence)

$$k(\theta_2|\theta_1) = \int_X p(\mathbf{x}|\theta_1) \log \frac{p(\mathbf{x}|\theta_1)}{p(\mathbf{x}|\theta_2)} d\mathbf{x}$$
(1)

- may diverge when the support of approximation and true distribution are different
- is not symmetric

Intrinsic discrepancy loss

$$\delta_X(\theta_1, \theta_2) = \min\{k(\theta_2|\theta_1), k(\theta_1|\theta_2)\}$$
(2)

- does not diverge even if the support of approximation and true distribution are different
- Symmetric, non-negative, invariant under one-to-one transformations
- does not represent local (lograrithmic form) score function

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- does not represent local (lograrithmic form) score function
- Intrinsic estimator

$$\theta^*(\mathbf{x}) = \arg\min_{\theta^e \in \Theta} \int_{\Theta} \delta(\theta^e, \theta) \pi_{\delta}(\theta | \mathbf{x}) d\theta$$
(3)

- $\pi_{\delta}(\theta | \mathbf{X}) d\theta$ is the reference posterior obtained using non-informative prior $\pi(\theta)$
 - Ensures (some kind of) objectivity

- For more information see:
 - Bernardo and Juárez. Intrinsic Estimation, Bayesian Statistics 7, 2003
 - http://www.uv.es/ bernardo/BernardoJuarez.pdf

Excersize

Let $\mathbf{x} = \{x_1, ..., x_n\}$, be a random sample from from the uniform distribution $\text{Un}(x|0, \theta) = \theta^{-1}, 0 < x < \theta$. Let $\text{Un}(x|0, \theta_1) = \theta_1^{-1}, 0 < x < \theta_1$ be the approximate distribution.

- What is the discrepancy of an approximation by definition 3.20 (Kullback-Leibler divergence)?
- What is the intrinsic discrepancy loss (2) of an approximation?
- How did the above results illustrate the behaviour of these different discrepancies in the case of different supports of an approximation and the true distribution?