

# Introduction to Expectation Propagation

**Antti Honkela**

Helsinki University of Technology

Espoo, Finland

<http://www.cis.hut.fi/ahonkela/>

# Contents

- Approximations and distance measures on distributions
- Limitations of naïve mean field variational Bayes (VB)
- Expectation propagation (EP) and the clutter problem
- Belief networks, loopy belief propagation and EP
- An energy function for EP

# Background

- Observations  $\mathcal{D}$ , model  $\mathcal{H}$  with parameters  $\theta$
- All information of the parameters is contained in the posterior

$$p(\theta|\mathcal{D}, \mathcal{H}) = \frac{p(\mathcal{D}|\theta, \mathcal{H})p(\theta|\mathcal{H})}{p(\mathcal{D}|\mathcal{H})},$$

where  $p(\mathcal{D}|\mathcal{H}) = \int_{\theta} p(\mathcal{D}|\theta, \mathcal{H})p(\theta|\mathcal{H})d\theta$

- Marginalisation principle:

$$p(\mathbf{x}|\mathcal{D}, \mathcal{H}) = \int_{\theta} p(\mathbf{x}|\theta, \mathcal{H})p(\theta|\mathcal{D}, \mathcal{H})d\theta$$

- How to assess possible approximations  $q(\theta)$  of the posterior  $p(\theta|\mathcal{D}, \mathcal{H})$ ?
- How to approximate  $p(\mathcal{D}|\mathcal{H})$ ?

# Bayesian analysis of approximations

- Choosing the best approximation is a decision problem
- Bayesian method: specify utility, maximise expected utility
- For approximations  $q(\boldsymbol{\theta}) \in \mathcal{Q}$  and “true parameter values”  $\boldsymbol{\theta} \in \Omega$ , define a **score function**  $u : \mathcal{Q} \times \Omega \rightarrow \mathbb{R}$
- Expected utility

$$\bar{u}(q) = \int u(q, \boldsymbol{\theta}) p(\boldsymbol{\theta} | \mathcal{D}) d\boldsymbol{\theta}'$$

## Properties of score functions

- The score function is **proper**, if

$$\sup \bar{u}(q) = \bar{u}(p(\boldsymbol{\theta}|\mathcal{D}))$$

which is attained only if  $q(\boldsymbol{\theta}) = p(\boldsymbol{\theta}|\mathcal{D})$

- The score function is **local**, if

$$u(q, \boldsymbol{\theta}) = u_{\boldsymbol{\theta}}(q(\boldsymbol{\theta}))$$

## Score functions

**Example.** The quadratic score function

$$u(q, \boldsymbol{\theta}) = A \left[ 2q(\boldsymbol{\theta}) - \int q(\boldsymbol{\theta}')^2 d\boldsymbol{\theta}' \right] + B(\boldsymbol{\theta})$$

corresponding to the expected utility

$$\bar{u}(q) = - \int (q(\boldsymbol{\theta}) - p(\boldsymbol{\theta}|\mathcal{D}))^2 d\boldsymbol{\theta}$$

is a **proper, non-local** score function

## Bayesian analysis of approximations

**Proposition.** Smooth, proper, local score functions are of the form

$$u(q, \boldsymbol{\theta}) = A \log q(\boldsymbol{\theta}) + B(\boldsymbol{\theta}),$$

where  $A > 0$  and  $B(\boldsymbol{\theta})$  are arbitrary.

**Proof.** We maximise the expected utility

$$\bar{u}(q) = \int u_{\boldsymbol{\theta}}(q(\boldsymbol{\theta}))p(\boldsymbol{\theta}|\mathcal{D})d\boldsymbol{\theta}$$

subject to constraint  $\int q(\boldsymbol{\theta})d\boldsymbol{\theta} = 1$ . This is done by finding an extremum of

$$F(q(\cdot)) = \bar{u}(q) - A \left[ \int q(\boldsymbol{\theta})d\boldsymbol{\theta} - 1 \right].$$

### Proof contd.

A necessary condition for this follows from the variational principle

$$\frac{\partial}{\partial \alpha} F(q(\cdot) + \alpha \tau(\cdot)) \Big|_{\alpha=0} = 0$$

for any function  $\tau : \Omega \rightarrow \mathbb{R}$ . this implies a differential equation

$$u'(q(\boldsymbol{\theta}))p(\boldsymbol{\theta}|\mathcal{D}) - A = 0,$$

which should hold for  $q(\boldsymbol{\theta}) = p(\boldsymbol{\theta}|\mathcal{D})$ . The solutions of this are

$$u(q, \boldsymbol{\theta}) = A \log q(\boldsymbol{\theta}) + B(\boldsymbol{\theta}).$$



## Bayesian analysis of approximations

**Theorem.** Differences of expected utilities under **smooth, proper, local** score functions are given by the (scaled) Kullback–Leibler (KL) divergence

$$A \cdot D_{KL}(p(\boldsymbol{\theta}|\mathcal{D}) \parallel q(\boldsymbol{\theta})) = A \int p(\boldsymbol{\theta}|\mathcal{D}) \log \frac{p(\boldsymbol{\theta}|\mathcal{D})}{q(\boldsymbol{\theta})} d\boldsymbol{\theta}.$$

**Proof.** Evaluate  $\bar{u}(p(\boldsymbol{\theta}|\mathcal{D})) - \bar{u}(q(\boldsymbol{\theta}))$ .

## Properties of KL divergence

- In information theory, the KL divergence

$$D_{KL}(p(\boldsymbol{\theta}|\mathcal{D}) || q(\boldsymbol{\theta})) = \int p(\boldsymbol{\theta}|\mathcal{D}) \log \frac{p(\boldsymbol{\theta}|\mathcal{D})}{q(\boldsymbol{\theta})} d\boldsymbol{\theta}$$

measures the overhead when using distribution  $q$  to code events following  $p$

- The choice of  $A$  reflects the choice of unit of measure, essentially the base of the logarithm
- Natural logarithm  $\ln$  yields **nats**, while  $\log_2$  gives **bits**

# Exponential families

**Definition** A set of distributions with densities

$$p(\boldsymbol{\theta}|\boldsymbol{\xi}) = \frac{1}{Z(\boldsymbol{\xi})} \exp(\boldsymbol{\xi}^T \phi(\boldsymbol{\theta}))$$

is an exponential family with **natural parameters**  $\boldsymbol{\xi}$ , **sufficient statistics**  $\phi(\boldsymbol{\theta})$  and **partition function**  $Z(\boldsymbol{\xi})$ .

Examples: Gaussian, gamma, multinomial, Dirichlet, ...

**Theorem** For exponential families,

$$\nabla_{\boldsymbol{\xi}} \log Z(\boldsymbol{\xi}) = \langle \phi(\boldsymbol{\theta}) \rangle.$$

## Properties of the KL divergence

**Theorem.** Given an approximation in an exponential family

$$q(\boldsymbol{\theta}|\boldsymbol{\xi}) = \frac{1}{Z(\boldsymbol{\xi})} \exp(\boldsymbol{\xi}^T \phi(\boldsymbol{\theta})),$$

the KL divergence  $D_{KL}(p(\boldsymbol{\theta}|\mathcal{D}) || q(\boldsymbol{\theta}|\boldsymbol{\xi}))$  is minimized when

$$\langle \phi(\boldsymbol{\theta}) \rangle_{p(\boldsymbol{\theta}|\mathcal{D})} = \langle \phi(\boldsymbol{\theta}) \rangle_{q(\boldsymbol{\theta}|\boldsymbol{\xi})}.$$

**Proof.** Consider

$$\begin{aligned} f(\boldsymbol{\xi}) &= D_{KL}(p(\boldsymbol{\theta}|\mathcal{D}) \parallel q(\boldsymbol{\theta}|\boldsymbol{\xi})) = \langle \log p \rangle_p + \langle \log Z(\boldsymbol{\xi}) \rangle_p - \langle \boldsymbol{\xi}^T \phi(\boldsymbol{\theta}) \rangle_p \\ &= \langle \log p \rangle_p + \log Z(\boldsymbol{\xi}) - \boldsymbol{\xi}^T \langle \phi(\boldsymbol{\theta}) \rangle_p. \end{aligned}$$

Zeroing the gradient yields the desired condition, because for exponential families

$$\nabla_{\boldsymbol{\xi}} \log Z(\boldsymbol{\xi}) = \langle \phi(\boldsymbol{\theta}) \rangle.$$

The minimality of the extremum can be checked using the second derivatives.

## Properties of the KL divergence

- In VB, the reverse of KL divergence is used:

$$D_{KL}(q(\boldsymbol{\theta}) \parallel p(\boldsymbol{\theta}|\mathcal{D})) = \int q(\boldsymbol{\theta}) \log \frac{q(\boldsymbol{\theta})}{p(\boldsymbol{\theta}|\mathcal{D})} d\boldsymbol{\theta}.$$

- Having large  $q(\boldsymbol{\theta})$  with very small  $p(\boldsymbol{\theta}|\mathcal{D})$  causes large values of the divergence
- Hence the VB approximation will be contained in the true distribution

## Limitations of naïve mean field variational Bayes

- The marginal likelihoods and especially rankings evaluated by VB are often quite reliable
- The estimates of the marginals may not be as good, variances can be underestimated
- Sometimes a simpler mode of solution may be preferred because of inadequate approximation

# Analysis of variational Bayesian ICA (A. Ilin & H. Valpola)

- Consider the ICA model

$$\mathbf{x} = \mathbf{A}\mathbf{s} + \mathbf{n}$$

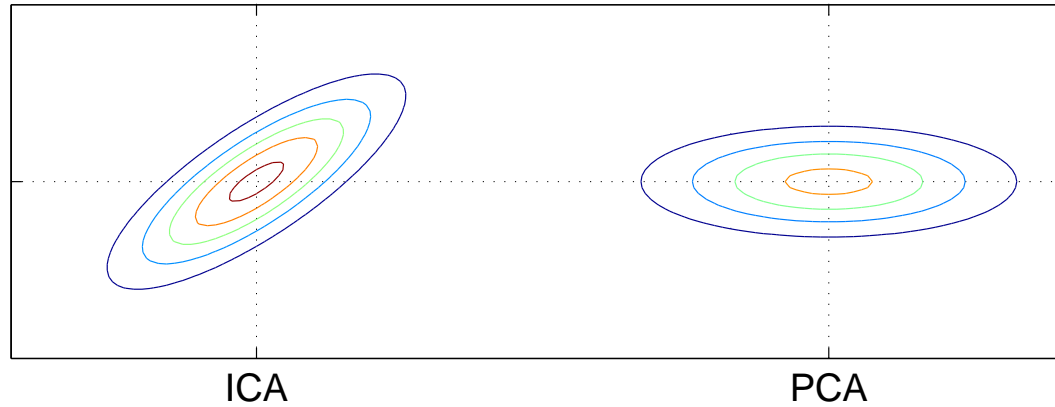
- Gaussian noise  $\mathbf{n} \sim \mathcal{N}(0, \Sigma_{\mathbf{x}})$
- Non-Gaussian source prior  $p(\mathbf{s}) = \prod_i p(s_i)$
- These yield non-diagonal posterior covariance for  $\mathbf{s}$ :

$$\Sigma_{\mathbf{s}|\mathcal{D}}^{-1} \propto \Sigma_{\mathbf{s}}^{-1} + \mathbf{A}^T \Sigma_{\mathbf{x}}^{-1} \mathbf{A}$$



# Limitations of variational Bayes

The form of the true posterior  $p(s(t) | A, x(t))$



The cost of the posterior and source model misfit

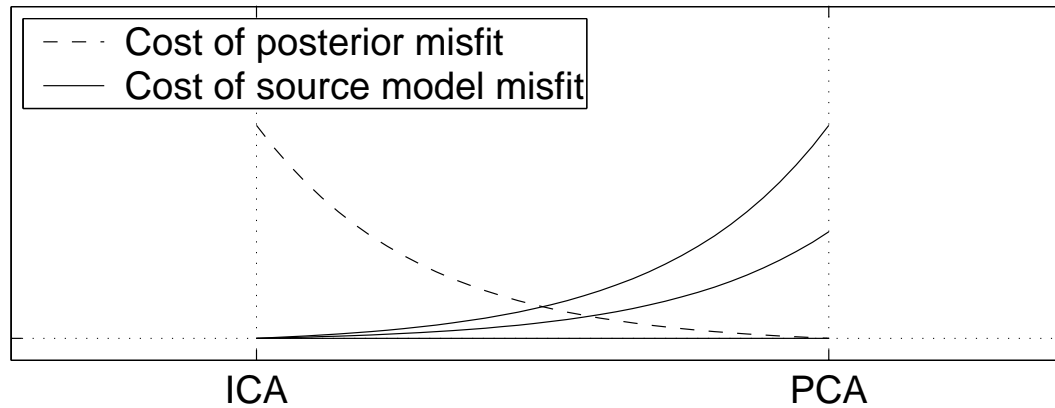
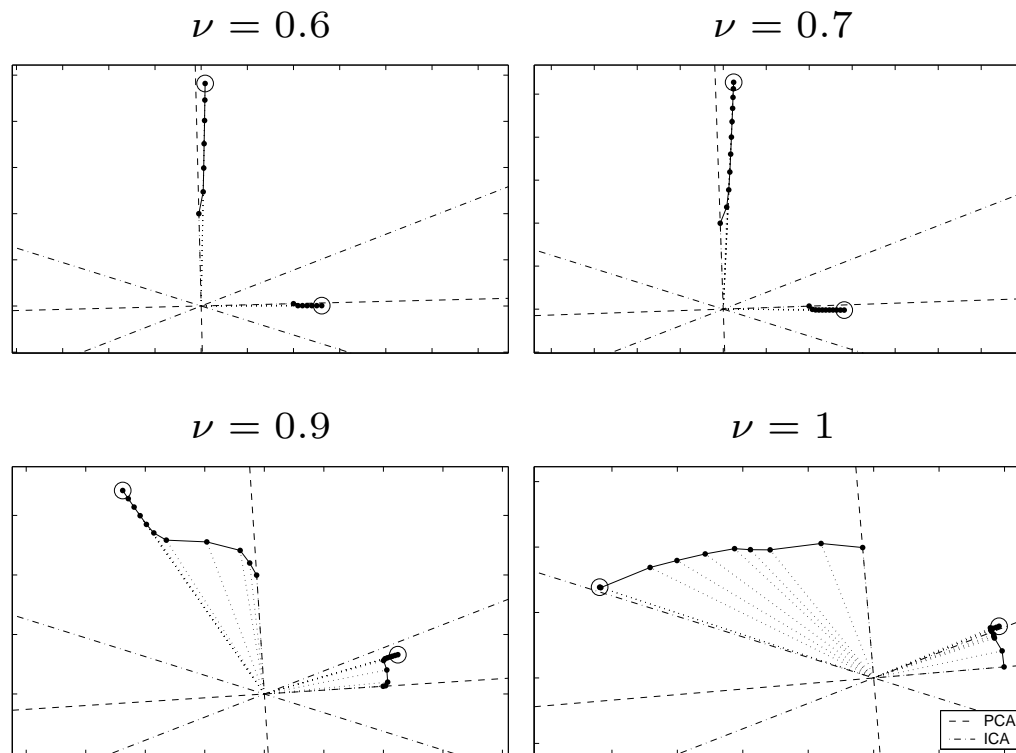


Illustration of the trade-offs between the ICA and PCA solutions.

# Limitations of variational Bayes



VB solutions to ICA problem as a function of non-Gaussianity of the sources

# Expectation propagation

- An approximate inference method proposed by Thomas Minka in 2001
- Suitable for approximating product forms

$$\prod_{i=0}^N t_i(\boldsymbol{\theta}) \approx \prod_{i=0}^N \tilde{t}_i(\boldsymbol{\theta})$$

- Iterative refinement of the terms  $\tilde{t}_i(\boldsymbol{\theta})$

# Expectation propagation

- The parameter posterior is

$$p(\boldsymbol{\theta}|\mathcal{D}) = \frac{1}{p(\mathcal{D})} p(\boldsymbol{\theta}) \prod_{i=1}^N p(\mathbf{x}_i|\boldsymbol{\theta})$$

- As a function of  $\boldsymbol{\theta}$ , this can be written as

$$p(\boldsymbol{\theta}) \prod_{i=1}^N p(\mathbf{x}_i|\boldsymbol{\theta}) = \prod_{i=0}^N t_i(\boldsymbol{\theta})$$

where  $t_0(\boldsymbol{\theta}) = p(\boldsymbol{\theta})$  and  $t_i(\boldsymbol{\theta}) = p(\mathbf{x}_i|\boldsymbol{\theta})$

- Now approximate each term separately to get

$$q(\boldsymbol{\theta}) = \prod_{i=0}^N \tilde{t}_i(\boldsymbol{\theta})$$

- Fit the approximation by finding

$$\min_{\tilde{t}_i(\boldsymbol{\theta})} D_{KL}(t_i(\boldsymbol{\theta}) \prod_{j \neq i} \tilde{t}_j(\boldsymbol{\theta}) \parallel \tilde{t}_i(\boldsymbol{\theta}) \prod_{j \neq i} \tilde{t}_j(\boldsymbol{\theta}))$$

# Expectation propagation algorithm

Input  $t_0(\boldsymbol{\theta}), \dots, t_N(\boldsymbol{\theta})$

Initialise  $\tilde{t}_0(\boldsymbol{\theta}) = t_0(\boldsymbol{\theta}), \tilde{t}_i(\boldsymbol{\theta}) = 1$  for  $i > 0$ ,  $q(\boldsymbol{\theta}) = \prod_{i=0}^N \tilde{t}_i(\boldsymbol{\theta})$

**repeat**

**for**  $i = 0, \dots, N$  **do**

**Deletion:**  $q_{\setminus i}(\boldsymbol{\theta}) \propto \frac{q(\boldsymbol{\theta})}{\tilde{t}_i(\boldsymbol{\theta})} = \prod_{j \neq i} \tilde{t}_j(\boldsymbol{\theta})$

**Projection:**  $\tilde{t}_i^{\text{new}}(\boldsymbol{\theta}) \leftarrow \arg \min_{\tilde{t}_i(\boldsymbol{\theta})} D_{KL}(t_i(\boldsymbol{\theta})q_{\setminus i}(\boldsymbol{\theta}) \parallel \tilde{t}_i(\boldsymbol{\theta})q_{\setminus i}(\boldsymbol{\theta}))$

**Inclusion:**  $q(\boldsymbol{\theta}) \leftarrow \tilde{t}_i^{\text{new}}(\boldsymbol{\theta})q_{\setminus i}(\boldsymbol{\theta})$

**end for**

**until** convergence

## Expectation propagation algorithm (2)

Input  $t_0(\boldsymbol{\theta}), \dots, t_N(\boldsymbol{\theta})$

Initialise  $\tilde{t}_0(\boldsymbol{\theta}) = t_0(\boldsymbol{\theta}), \tilde{t}_i(\boldsymbol{\theta}) = 1$  for  $i > 0$ ,  $q(\boldsymbol{\theta}) = \prod_{i=0}^N \tilde{t}_i(\boldsymbol{\theta})$

**repeat**

**for**  $i = 0, \dots, N$  **do**

**Deletion:**  $q_{\setminus i}(\boldsymbol{\theta}) \propto \frac{q(\boldsymbol{\theta})}{\tilde{t}_i(\boldsymbol{\theta})} = \prod_{j \neq i} \tilde{t}_j(\boldsymbol{\theta})$

**Inclusion:**  $q(\boldsymbol{\theta}) \leftarrow \arg \min_{q(\boldsymbol{\theta})} D_{KL}(t_i(\boldsymbol{\theta})q_{\setminus i}(\boldsymbol{\theta}) \parallel q(\boldsymbol{\theta}))$

**Update:**  $\tilde{t}_i^{\text{new}}(\boldsymbol{\theta}) \leftarrow \frac{q(\boldsymbol{\theta})}{q_{\setminus i}(\boldsymbol{\theta})}$

**end for**

**until** convergence

## The clutter problem

Consider a simple Gaussian mixture for  $\mathcal{D} = \{\mathbf{x}_i\}_{i=1}^N$

$$p(\mathbf{x}|\boldsymbol{\theta}) = w\mathcal{N}(\mathbf{x}; \boldsymbol{\theta}, \mathbf{I}) + (1 - w)\mathcal{N}(\mathbf{x}; \mathbf{0}, 10\mathbf{I})$$

$$p(\boldsymbol{\theta}) = \mathcal{N}(\boldsymbol{\theta}; \mathbf{0}, 100\mathbf{I}).$$

A suitable exponential family for this is formed by

$$\mathcal{N}(\mathbf{x}; \mathbf{m}, v\mathbf{I}) = \mathcal{N}(\mathbf{x}; \boldsymbol{\xi})$$

with sufficient statistics  $\phi(\mathbf{x}) = (\mathbf{x}, \mathbf{x}^T \mathbf{x})$ , natural parameters

$\boldsymbol{\xi} = (v^{-1}\mathbf{m}, -\frac{1}{2}v^{-1})$  and normalisation  $Z(\boldsymbol{\xi}) = (2\pi v)^{d/2} \exp(\frac{1}{2v}\mathbf{m}^T \mathbf{m})$ .



# Expectation propagation algorithm

Input  $t_0(\boldsymbol{\theta}), \dots, t_N(\boldsymbol{\theta})$

Initialise  $\tilde{t}_0(\boldsymbol{\theta}) = t_0(\boldsymbol{\theta}), \tilde{t}_i(\boldsymbol{\theta}) = 1$  for  $i > 0$ ,  $q(\boldsymbol{\theta}) = \prod_{i=0}^N \tilde{t}_i(\boldsymbol{\theta})$

repeat

  for  $i = 0, \dots, N$  do

    Deletion:  $q_{\setminus i}(\boldsymbol{\theta}) \propto \frac{q(\boldsymbol{\theta})}{\tilde{t}_i(\boldsymbol{\theta})} = \prod_{j \neq i} \tilde{t}_j(\boldsymbol{\theta})$

    Inclusion:  $q(\boldsymbol{\theta}) \leftarrow \arg \min_{q(\boldsymbol{\theta})} D_{KL}(t_i(\boldsymbol{\theta})q_{\setminus i}(\boldsymbol{\theta}) \parallel q(\boldsymbol{\theta}))$

    Update:  $\tilde{t}_i^{\text{new}}(\boldsymbol{\theta}) \leftarrow \frac{q(\boldsymbol{\theta})}{q_{\setminus i}(\boldsymbol{\theta})}$

  end for

until convergence

# EP for the clutter problem (1): Initialisation

For the clutter problem, we have

$$t_0(\boldsymbol{\theta}) = p(\boldsymbol{\theta})$$

$$t_i(\boldsymbol{\theta}) = p(\mathbf{x}_i|\boldsymbol{\theta}), \quad i = 1, \dots, N.$$

The approximation is of the form

$$\tilde{t}_0(\boldsymbol{\theta}) = t_0(\boldsymbol{\theta}) = p(\boldsymbol{\theta})$$

$$\tilde{t}_i(\boldsymbol{\theta}) = s_i \exp(\boldsymbol{\xi}_i^T \phi(\boldsymbol{\theta})), \quad i = 1, \dots, N,$$

$$q(\boldsymbol{\theta}) = \prod_{i=0}^N \tilde{t}_i(\boldsymbol{\theta}) = s \mathcal{N}(\boldsymbol{\theta}; \boldsymbol{\xi})$$

Now initialise  $\boldsymbol{\xi}_i = \mathbf{0}$  for  $i = 1, \dots, N$ .

# Expectation propagation algorithm

Input  $t_0(\boldsymbol{\theta}), \dots, t_N(\boldsymbol{\theta})$

Initialise  $\tilde{t}_0(\boldsymbol{\theta}) = t_0(\boldsymbol{\theta}), \tilde{t}_i(\boldsymbol{\theta}) = 1$  for  $i > 0$ ,  $q(\boldsymbol{\theta}) = \prod_{i=0}^N \tilde{t}_i(\boldsymbol{\theta})$

repeat

  for  $i = 0, \dots, N$  do

    Deletion:  $q_{\setminus i}(\boldsymbol{\theta}) \propto \frac{q(\boldsymbol{\theta})}{\tilde{t}_i(\boldsymbol{\theta})} = \prod_{j \neq i} \tilde{t}_j(\boldsymbol{\theta})$

    Inclusion:  $q(\boldsymbol{\theta}) \leftarrow \arg \min_{q(\boldsymbol{\theta})} D_{KL}(t_i(\boldsymbol{\theta})q_{\setminus i}(\boldsymbol{\theta}) || q(\boldsymbol{\theta}))$

    Update:  $\tilde{t}_i^{\text{new}}(\boldsymbol{\theta}) \leftarrow \frac{q(\boldsymbol{\theta})}{q_{\setminus i}(\boldsymbol{\theta})}$

  end for

until convergence

## EP for the clutter problem (2): Deletion

When working with natural parameters, the deletion operation

$$q_{\setminus i}(\boldsymbol{\theta}) \propto \frac{q(\boldsymbol{\theta})}{\tilde{t}_i(\boldsymbol{\theta})}$$

is trivial to implement with

$$\boldsymbol{\xi}_{\setminus i} = \boldsymbol{\xi} - \boldsymbol{\xi}_i.$$

# Expectation propagation algorithm

Input  $t_0(\boldsymbol{\theta}), \dots, t_N(\boldsymbol{\theta})$

Initialise  $\tilde{t}_0(\boldsymbol{\theta}) = t_0(\boldsymbol{\theta}), \tilde{t}_i(\boldsymbol{\theta}) = 1$  for  $i > 0$ ,  $q(\boldsymbol{\theta}) = \prod_{i=0}^N \tilde{t}_i(\boldsymbol{\theta})$

repeat

  for  $i = 0, \dots, N$  do

    Deletion:  $q_{\setminus i}(\boldsymbol{\theta}) \propto \frac{q(\boldsymbol{\theta})}{\tilde{t}_i(\boldsymbol{\theta})} = \prod_{j \neq i} \tilde{t}_j(\boldsymbol{\theta})$

    Inclusion:  $q(\boldsymbol{\theta}) \leftarrow \arg \min_{q(\boldsymbol{\theta})} D_{KL}(t_i(\boldsymbol{\theta})q_{\setminus i}(\boldsymbol{\theta}) || q(\boldsymbol{\theta}))$

    Update:  $\tilde{t}_i^{\text{new}}(\boldsymbol{\theta}) \leftarrow \frac{q(\boldsymbol{\theta})}{q_{\setminus i}(\boldsymbol{\theta})}$

  end for

until convergence

## EP for the clutter problem (3): Inclusion

The inclusion operation:

$$q(\boldsymbol{\theta}) \leftarrow \arg \min_{q(\boldsymbol{\theta})} D_{KL}(t_i(\boldsymbol{\theta})q_{\setminus i}(\boldsymbol{\theta}) \parallel q(\boldsymbol{\theta}))$$

requires matching sufficient statistics of

$$\begin{aligned} t_i(\boldsymbol{\theta})q_{\setminus i}(\boldsymbol{\theta}) &= (w\mathcal{N}(\mathbf{x}_i; \boldsymbol{\theta}, \mathbf{I}) + (1-w)\mathcal{N}(\mathbf{x}_i; \mathbf{0}, 10\mathbf{I}))\mathcal{N}(\boldsymbol{\theta}; \boldsymbol{\xi}_{\setminus i}) \\ &= \left( w\mathcal{N}\left(\boldsymbol{\theta}; \left(\mathbf{x}_i, -\frac{1}{2}\right)\right) + (1-w)\mathcal{N}(\mathbf{x}_i; \mathbf{0}, 10\mathbf{I}) \right) \mathcal{N}(\boldsymbol{\theta}; \boldsymbol{\xi}_{\setminus i}) \\ &= w \frac{Z(\boldsymbol{\xi}^+)}{Z\left(\left(\mathbf{x}_i, -\frac{1}{2}\right)\right) Z(\boldsymbol{\xi}_{\setminus i})} \mathcal{N}(\boldsymbol{\theta}; \boldsymbol{\xi}^+) + (1-w)\mathcal{N}(\mathbf{x}_i; \mathbf{0}, 10\mathbf{I})\mathcal{N}(\boldsymbol{\theta}; \boldsymbol{\xi}_{\setminus i}) \\ &\propto r\mathcal{N}(\boldsymbol{\theta}; \boldsymbol{\xi}^+) + (1-r)\mathcal{N}(\boldsymbol{\theta}; \boldsymbol{\xi}_{\setminus i}), \end{aligned}$$

where  $\boldsymbol{\xi}^+ = \boldsymbol{\xi}_{\setminus i} + \left(\mathbf{x}_i, -\frac{1}{2}\right)$

## EP for the clutter problem (3): Inclusion (cont.)

We wish to match the sufficient statistics of the Gaussian mixture

$$t_i(\boldsymbol{\theta})q_{\setminus i}(\boldsymbol{\theta}) \propto r\mathcal{N}(\boldsymbol{\theta}; \boldsymbol{\xi}^+) + (1-r)\mathcal{N}(\boldsymbol{\theta}; \boldsymbol{\xi}_{\setminus i}).$$

These are simply

$$\mathbf{m} = r\mathbf{m}^+ + (1-r)\mathbf{m}_{\setminus i}$$
$$v + \mathbf{m}^T \mathbf{m} = r(v^+ + (\mathbf{m}^+)^T \mathbf{m}^+) + (1-r)(v_{\setminus i} + \mathbf{m}_{\setminus i}^T \mathbf{m}_{\setminus i})$$

# Expectation propagation algorithm

Input  $t_0(\boldsymbol{\theta}), \dots, t_N(\boldsymbol{\theta})$

Initialise  $\tilde{t}_0(\boldsymbol{\theta}) = t_0(\boldsymbol{\theta}), \tilde{t}_i(\boldsymbol{\theta}) = 1$  for  $i > 0$ ,  $q(\boldsymbol{\theta}) = \prod_{i=0}^N \tilde{t}_i(\boldsymbol{\theta})$

**repeat**

**for**  $i = 0, \dots, N$  **do**

Deletion:  $q_{\setminus i}(\boldsymbol{\theta}) \propto \frac{q(\boldsymbol{\theta})}{\tilde{t}_i(\boldsymbol{\theta})} = \prod_{j \neq i} \tilde{t}_j(\boldsymbol{\theta})$

Inclusion:  $q(\boldsymbol{\theta}) \leftarrow \arg \min_{q(\boldsymbol{\theta})} D_{KL}(t_i(\boldsymbol{\theta})q_{\setminus i}(\boldsymbol{\theta}) \parallel q(\boldsymbol{\theta}))$

Update:  $\tilde{t}_i^{\text{new}}(\boldsymbol{\theta}) \leftarrow \frac{q(\boldsymbol{\theta})}{q_{\setminus i}(\boldsymbol{\theta})}$

**end for**

**until** convergence



## EP for the clutter problem (4): Update

When working with natural parameters, the update operation

$$\tilde{t}_i^{\text{new}}(\boldsymbol{\theta}) \leftarrow \frac{q(\boldsymbol{\theta})}{q_{\setminus i}(\boldsymbol{\theta})}$$

is again trivial with

$$\boldsymbol{\xi}_i = \boldsymbol{\xi} - \boldsymbol{\xi}_{\setminus i}.$$

## Marginal likelihood by EP

- The EP algorithm may be extended to evaluate the marginal likelihood  $p(\mathcal{D}|\mathcal{H})$
- To do this, we include a scale on  $\tilde{t}_i(\boldsymbol{\theta})$  and through them for  $q(\boldsymbol{\theta})$ :

$$\tilde{t}_i(\boldsymbol{\theta}) = Z_i \frac{q^*(\boldsymbol{\theta})}{q_{\setminus i}(\boldsymbol{\theta})},$$

where  $q^*(\boldsymbol{\theta})$  is a normalised version of  $q(\boldsymbol{\theta})$  and

$$Z_i = \int q_{\setminus i}(\boldsymbol{\theta}) t_i(\boldsymbol{\theta}) d\boldsymbol{\theta}$$

- Finally we approximate

$$p(\mathcal{D}|\mathcal{H}) \approx \int q(\boldsymbol{\theta}) d\boldsymbol{\theta} = \int \prod_i \tilde{t}_i(\boldsymbol{\theta}) d\boldsymbol{\theta}$$

# Marginal likelihood for the clutter problem

For the clutter problem

$$\tilde{t}_i(\boldsymbol{\theta}) = Z_i \frac{q^*(\boldsymbol{\theta})}{q_{\setminus i}(\boldsymbol{\theta})}$$

implies

$$s_i = Z_i \frac{Z(\boldsymbol{\xi}_{\setminus i})}{Z(\boldsymbol{\xi})}$$

$$Z_i = w \frac{Z(\boldsymbol{\xi}^+)}{Z\left(\left(\mathbf{x}_i, -\frac{1}{2}\right)\right) Z(\boldsymbol{\xi}_{\setminus i})} + (1-w)\mathcal{N}(\mathbf{x}_i; \mathbf{0}, 10\mathbf{I}).$$

And globally

$$p(\mathcal{D}|\mathcal{H}) \approx \int \prod_i \tilde{t}_i(\boldsymbol{\theta}) d\boldsymbol{\theta} = \frac{Z(\boldsymbol{\xi})}{Z(\boldsymbol{\xi}_0)} \prod_{i=1}^N s_i$$

## EP for belief networks

- A probabilistic model may be represented as a directed graph corresponding to a factorisation of the joint distribution

$$p(\mathbf{x}) = \prod_{x_i \in \mathbf{x}} p(x_i | \text{parents}(x_i))$$

- Derive an EP algorithm using the term factorisation

$$t_i(\mathbf{x}) = p(x_i | \text{parents}(x_i))$$

and a factorial posterior approximation

$$q(\mathbf{x}) = \prod_k q_k(x_k)$$

- For each term  $t_i(\mathbf{x})$  the factorisation implies a factorial approximation

$$\tilde{t}_i(\mathbf{x}) = \prod_{k \in \{i, \text{pa}(i)\}} \tilde{t}_{ik}(x_k)$$

- Equivalently, for each factor  $q_k(x_k)$ , this corresponds to a regular EP approximation

$$q_k(x_k) = \prod_{i \in \{i, \text{ch}(i)\}} \tilde{t}_{ik}(x_k),$$

## EP for belief networks

Input  $t_1(\mathbf{x}), \dots, t_N(\mathbf{x})$

Initialise  $\tilde{t}_{ik}(x_k) = 1, q_k(x_k) = \prod_i \tilde{t}_{ik}(x_k)$

**repeat**

**for**  $i = 1, \dots, N$  **do**

**for all**  $k$  **do**

**Deletion:**  $q_{\setminus i,k}(x_k) \propto \frac{q_k(x_k)}{\tilde{t}_{ik}(x_k)} = \prod_{j \neq i} \tilde{t}_{jk}(x_k)$

**end for**

**for all**  $k$  **do**

**Projection:**  $\tilde{t}_{ik}^{\text{new}}(x_k) \leftarrow \sum_{\mathbf{x} \setminus x_k} t_i(\mathbf{x}) \prod_{j \neq k} q_{\setminus i,j}(x_j)$

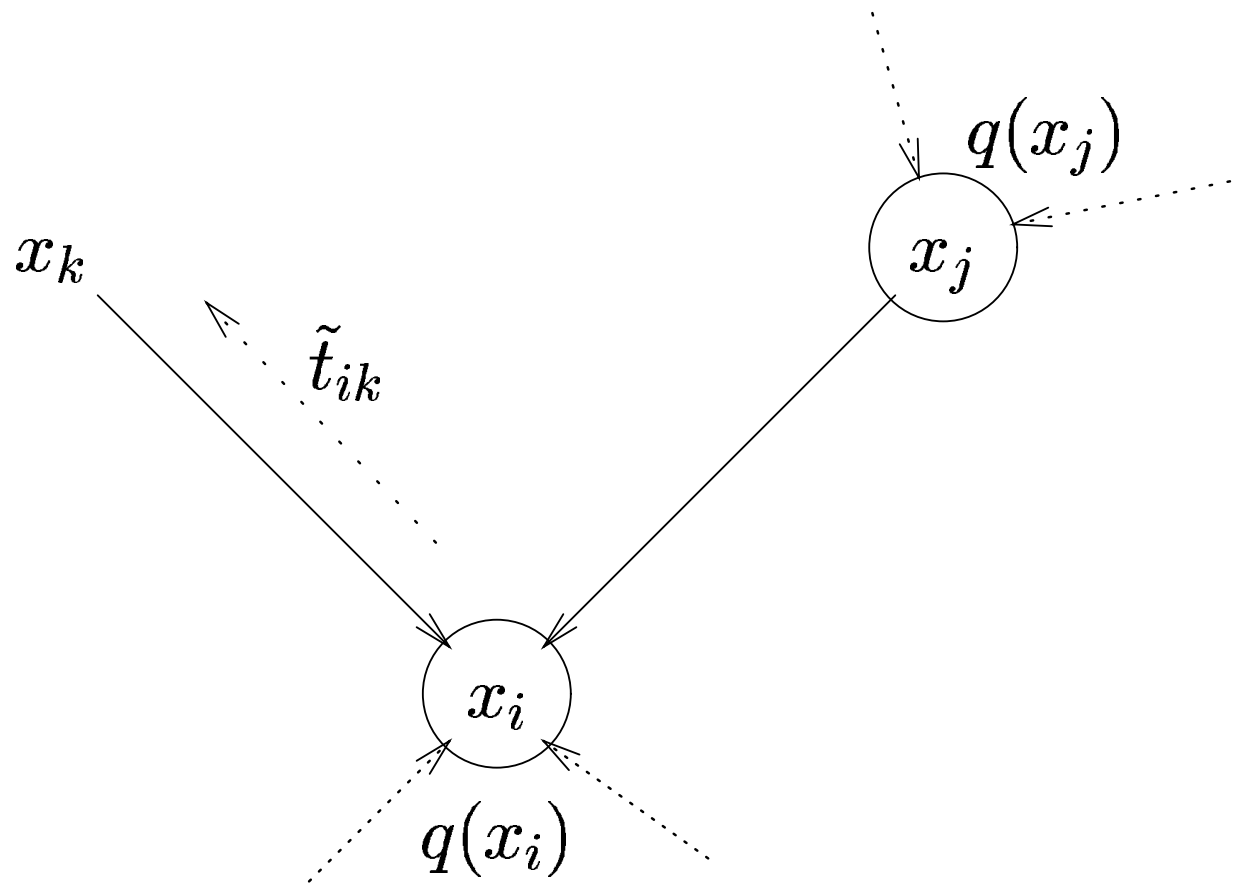
**Inclusion:**  $q_k(x_k) \leftarrow \tilde{t}_{ik}^{\text{new}}(x_k) q_{\setminus i,k}(x_k)$

**end for**

**end for**

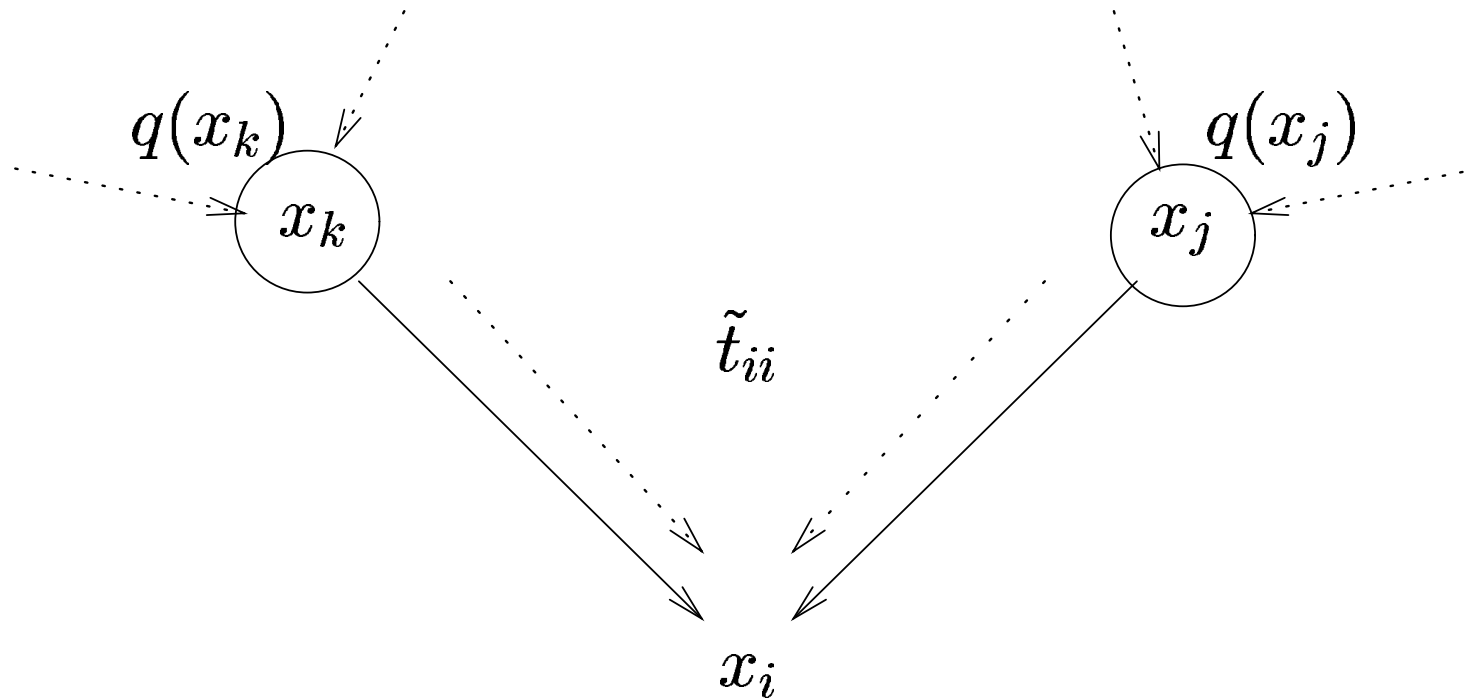
**until** convergence

## EP for belief networks (T. Minka)



$$\tilde{t}_{ik}(x_k) = \sum_{x_i, x_j} p(x_i | x_k, x_j) q_i(x_i) q_j(x_j)$$

## EP for belief networks (T. Minka)



$$\tilde{t}_{ii}(x_i) = \sum_{x_k, x_j} p(x_i | x_k, x_j) q_k(x_k) q_j(x_j)$$



## EP for belief networks

- The presented EP algorithm is equivalent to a well-known method called (loopy) belief propagation
- For tree structured graphs, it converges in one pass to yield correct marginals
- For general graphs there are no guarantees and it may even diverge

## EP for belief networks

- The EP formulation allows simple generalisation to more accurate approximations
- Use fewer more complicated terms  $t_i(\mathbf{x})$
- Factorisation  $q(\mathbf{x}) = \prod_k q_k(x_k)$  over nodes can still be assumed to only evaluate the marginals

## An energy function for EP

- Assume an approximation in an exponential family  $\exp(\boldsymbol{\lambda}^T \phi(\boldsymbol{\theta}))$
- With an exact prior,

$$q(\boldsymbol{\theta}) \propto p(\boldsymbol{\theta}) \exp(\boldsymbol{\nu}^T \phi(\boldsymbol{\theta}))$$

and

$$q_{\setminus i}(\boldsymbol{\theta}) = p(\boldsymbol{\theta}) \exp(\boldsymbol{\lambda}_i^T \phi(\boldsymbol{\theta}))$$

- Let  $N$  be the number of terms  $t_i(\boldsymbol{\theta})$

- Now, EP fixed points correspond to stationary points of the objective

$$\min_{\nu} \max_{\lambda} (N - 1) \log \int p(\boldsymbol{\theta}) \exp(\boldsymbol{\nu}^T \phi(\boldsymbol{\theta})) d\boldsymbol{\theta} - \sum_{i=1}^N \log \int t_i(\boldsymbol{\theta}) p(\boldsymbol{\theta}) \exp(\boldsymbol{\lambda}_i^T \phi(\boldsymbol{\theta})) d\boldsymbol{\theta}$$

such that  $(N - 1)\nu_j = \sum_i \lambda_{ij}$ .

- Note: non-convex optimisation problem
- Also other formulations for the energy function

## Summary

- Kullback–Leibler divergence  $D_{KL}(p(\boldsymbol{\theta}|\mathcal{D}) \parallel q(\boldsymbol{\theta}))$  is a reasonable measure of goodness of approximation
- EP uses this in a tractable manner to optimise

$$D_{KL}(t_i(\boldsymbol{\theta})q_{\setminus i}(\boldsymbol{\theta}) \parallel \tilde{t}_i(\boldsymbol{\theta})q_{\setminus i}(\boldsymbol{\theta}))$$

- Provides good approximations of marginals and marginal likelihood
- Alternative interpretation to existing belief net algorithms
- Algorithm may not converge ( $\rightarrow$  explicitly minimise the energy?)